

Markovian Queues and Stochastic Networks

Lecture 4 Richard J. Boucherie Stochastic Operations Research





Time-reversed process and Kelly's Lemma – 1

- Stationary Markov chain $\{N(t)\}$
- Time-reversed process $\{N^r(t)\} = \{N(\tau t)\}.$

$$\mathbb{P}(N(t) = \mathbf{n}|N(t+h) = \mathbf{n}') = \frac{\mathbb{P}(N(t) = \mathbf{n})}{\mathbb{P}(N(t+h) = \mathbf{n}')}\mathbb{P}(N(t+h) = \mathbf{n}'|N(t) = \mathbf{n}).$$

Theorem (4.1.2)

Let { $N(t), t \in \mathbb{R}$ } be stationary Markov chain with transition rates $q(\mathbf{n}, \mathbf{n}'), \mathbf{n}, \mathbf{n}' \in S$, and equilibrium distribution $\pi(\mathbf{n})$, $\mathbf{n} \in S$. The time-reversed process { $N(\tau - t), t \in \mathbb{R}$ } is a conservative, stable, regular, irreducible continuous-time stationary Markov chain with transition rates $q^r(\mathbf{n}, \mathbf{n}'), \mathbf{n}, \mathbf{n}' \in S$ given by

$$q^{\prime}(\mathbf{n},\mathbf{n}^{\prime})=rac{\pi(\mathbf{n}^{\prime})}{\pi(\mathbf{n})}q(\mathbf{n}^{\prime},\mathbf{n}),$$

and the same equilibrium distribution $\pi(\mathbf{n})$, $\mathbf{n} \in S$. UNIVERSITY OF TWENTE. Markovian Queues and Stochastic Networks Theorem (4.1.3 Kelly's lemma)

Let {N(t), $t \in \mathbb{R}$ } be a stationary Markov chain with transition rates $q(\mathbf{n}, \mathbf{n}')$, $\mathbf{n}, \mathbf{n}' \in S$. If we can find a collection of numbers $q'(\mathbf{n}, \mathbf{n}')$, $\mathbf{n}, \mathbf{n}' \in S$, such that

$$\sum_{\mathbf{n}'
eq \mathbf{n}} q(\mathbf{n},\mathbf{n}') = \sum_{\mathbf{n}'
eq \mathbf{n}} q'(\mathbf{n},\mathbf{n}'), \quad \mathbf{n}\in\mathcal{S},$$

and a distribution $\pi = (\pi(\mathbf{n}), \mathbf{n} \in S)$ such that

$$\pi(\mathbf{n})q'(\mathbf{n},\mathbf{n}')=\pi(\mathbf{n}')q(\mathbf{n}',\mathbf{n}),\quad \mathbf{n},\mathbf{n}'\in\mathcal{S},$$

then $q^r(\mathbf{n}, \mathbf{n}')$, $\mathbf{n}, \mathbf{n}' \in S$, are the transition rates of the time-reversed Markov chain $\{N(\tau - t), t \in \mathbb{R}\}$ and $\pi(\mathbf{n}), \mathbf{n} \in S$, is the equilibrium distribution of both Markov chains.

- Poisson arrival rate λ and service rate μ , $\rho = \lambda/\mu < 1$.
- Departure rate λ .
- Guess for arrival rate time-reversed process: $q^r(\mathbf{n}, \mathbf{n} + 1) = \lambda$.
- Further guess could then be $q^r(\mathbf{n}, \mathbf{n} 1) = \mu$
- Educated guess for the equilibrium distribution is $\pi(\mathbf{n}) = (1 \rho)\rho^{\mathbf{n}}$,
- Kelly's lemma 2 is satisfied.
- Time-reversed process is an *M*|*M*|1 queue with Poisson arrivals at rate λ and service rate μ.

Example: Kelly-Whittle network – 1

► Transition rates Kelly-Whittle network:

$$q(\mathbf{n},\mathbf{n}') = \begin{cases} \frac{\psi(\mathbf{n}-\mathbf{e}_i)}{\phi(\mathbf{n})}\mu_i\rho_{ij}, & \text{if } \mathbf{n}' = \mathbf{n}-\mathbf{e}_i+\mathbf{e}_j, \ i,j=0,\ldots,J, \\ 0, & \text{otherwise.} \end{cases}$$

 Kelly's lemma: the time-reversed routing process is the Markov chain with transition probabilities

$$p_{ij}^r := rac{\lambda_j}{\lambda_j} p_{ji}, \quad i, j = 0, \dots, J.$$

Natural guess for the transition rates of the time-reversed process {N^r(t)} is, for n ≠ n',

$$q^{r}(\mathbf{n},\mathbf{n}') = \begin{cases} \frac{\psi(\mathbf{n}-\mathbf{e}_{i})}{\phi(\mathbf{n})}\mu_{i}\rho_{ij}^{r}, & \text{if } \mathbf{n}'=\mathbf{n}-\mathbf{e}_{i}+\mathbf{e}_{j}, \ i,j=0,\ldots,J, \\ 0, & \text{otherwise.} \end{cases}$$

Example: Kelly-Whittle network - 2

Observe that

$$\sum_{\mathbf{n}'\neq\mathbf{n}} q(\mathbf{n},\mathbf{n}') = \sum_{i,j=0}^{J} \frac{\psi(\mathbf{n}-\mathbf{e}_i)}{\phi(\mathbf{n})} \mu_i p_{ij} = \sum_{i=0}^{J} \frac{\psi(\mathbf{n}-\mathbf{e}_i)}{\phi(\mathbf{n})} \mu_i,$$

$$\sum_{\mathbf{n}'\neq\mathbf{n}} q^r(\mathbf{n},\mathbf{n}') = \sum_{i,j=0}^{J} \frac{\psi(\mathbf{n}-\mathbf{e}_i)}{\phi(\mathbf{n})} \mu_i p_{ij}^r = \sum_{i,j=0}^{J} \frac{\psi(\mathbf{n}-\mathbf{e}_i)}{\phi(\mathbf{n})} \mu_i \frac{\lambda_j}{\lambda_i} p_{ji} = \sum_{i=0}^{J} \frac{\psi(\mathbf{n}-\mathbf{e}_i)}{\phi(\mathbf{n})} \mu_i,$$

► Educated guess: $\pi(\mathbf{n}) = G_{KW}\phi(\mathbf{n})\prod_{j=1}^{J}\rho_{j}^{n_{j}}, \mathbf{n} \in S$,

Then

$$\pi(\mathbf{n})q^{r}(\mathbf{n},\mathbf{n}') = G_{KW}\phi(\mathbf{n})\prod_{k=1}^{J}\rho_{k}^{n_{k}}\frac{\psi(\mathbf{n}-\mathbf{e}_{i})}{\phi(\mathbf{n})}\mu_{i}\frac{\lambda_{j}}{\lambda_{i}}\rho_{ji}$$

$$= G_{KW}\phi(\mathbf{n}-\mathbf{e}_{i}+\mathbf{e}_{j})\prod_{k=1}^{J}\rho_{k}^{n_{k}-\delta_{ki}+\delta_{kj}}\frac{\psi(\mathbf{n}-\mathbf{e}_{i})}{\phi(\mathbf{n}-\mathbf{e}_{i}+\mathbf{e}_{j})}\mu_{j}\rho_{ji}$$

$$= \pi(\mathbf{n}-\mathbf{e}_{i}+\mathbf{e}_{j})q(\mathbf{n}-\mathbf{e}_{i}+\mathbf{e}_{j},\mathbf{n}).$$

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Queue disciplines - 1

Definition (4.2.1 (κ, γ, δ)-protocol)

Customers ordered: if queue contains *n* customers then customers in positions $1, \ldots, n, n \in \mathbb{N}$. Queue operation:

- ► a customer of class c requires a negative-exponentially distributed amount of service with rate µ(c);
- if n > 0 customers present service at rate $\kappa(n) > 0$;
- Fraction γ(ℓ, n) of service to customer in position ℓ; if customer in position ℓ completes service then customers in positions ℓ + 1, ℓ + 2,..., n move to ℓ, ℓ + 1,..., n − 1;
- ► arriving customer into position ℓ with probability δ(ℓ, n + 1); customers in positions ℓ, ℓ + 1, ..., n move to positions ℓ + 1, ℓ + 2, ..., n + 1.

$$\sum_{\substack{\ell=1\\ \ell=1}}^{n} \gamma(\ell, n) = 1, \quad \sum_{\substack{\ell=1\\ \ell=1}}^{n} \delta(\ell, n) = 1.$$
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	κ(n)	$\gamma(\ell, n)$	$\delta(\ell, n)$
FIFO	1(n > 0)	$1(\ell = 1)$	$\mathbb{1}(\ell = n)$
LIFO-PR	1(n > 0)	$1(\ell = n)$	$\mathbb{1}(\ell = n)$
PS	1(n > 0)	1/n	1/n
INF	n	1/ <i>n</i>	1/ <i>n</i>

Markovian Queues and Stochastic Networks

Multi-class LIFO-PR queue - 1

- ► Customers of class c arrive according to a Poisson process with rate λ(c), ρ(c) = λ(c)/μ(c) < 1,</p>
- Let c = (c(1),..., c(n)) record the class of the customers in position i, i = 1,..., n,
- Let $\{N(t)\}$ record the state of the Markov chain at

$$S = \{ \mathbf{c} : \mathbf{c} = (c(1), \dots, c(n)), \ c(i) \in \{1, \dots, C\} \}.$$

Transition rates

$$q(\mathbf{c},\mathbf{c}') = \left\{egin{array}{ll} \lambda(m{c}), & ext{if } \mathbf{c}' = (m{c}(1),\ldots,m{c}(n),m{c}), \ \mu(m{c}(n)), & ext{if } \mathbf{c}' = (m{c}(1),\ldots,m{c}(n-1)). \end{array}
ight.$$

Natural guess for time-reversed queue is the queue multi-class LIFO-PR queue with the same rates:

$$q^{r}(\mathbf{c},\mathbf{c}') = \begin{cases} \lambda(\mathbf{c}), & \text{if } \mathbf{c}' = (\mathbf{c}(1),\ldots,\mathbf{c}(n),\mathbf{c}), \\ \mu(\mathbf{c}(n)), & \text{if } \mathbf{c}' = ((\mathbf{c}(1),\ldots,\mathbf{c}(n-1)). \end{cases}$$

Guess equilibrium distribution:

$$\pi(\mathbf{c}) = G \prod_{i=1}^{n} \rho(\mathbf{c}(i)), \quad \mathbf{c} = (\mathbf{c}(1), \dots, \mathbf{c}(n)) \in S.$$

Check conditions Kelly's lemma:

$$\sum_{\mathbf{c}'\neq\mathbf{c}} q(\mathbf{c},\mathbf{c}') = \sum_{c=1}^{C} \lambda(c) + \mu(c(n)) = \sum_{\mathbf{c}'\neq\mathbf{c}} q^{r}(\mathbf{c},\mathbf{c}').$$

For $\mathbf{c} = (c(1), \dots, c(n)), \mathbf{c}' = (c(1), \dots, c(n), c),$
 $\mathbf{c}'' = (c(1), \dots, c(n-1)),$
 $\pi(\mathbf{c})q^{r}(\mathbf{c},\mathbf{c}') = \pi(\mathbf{c}')q(\mathbf{c}',\mathbf{c}) \Leftrightarrow \lambda(c) = \rho(c)\mu(c),$
 $\pi(\mathbf{c})q^{r}(\mathbf{c},\mathbf{c}'') = \pi(\mathbf{c}'')q(\mathbf{c}'',\mathbf{c}) \Leftrightarrow \rho(c(n))\mu(c(n)) = \lambda(c(n))$



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Networks: customer types and fixed routes - 1

- ► Network of *J* queues.
- ► Customers of types u = 1,..., U, arrive to a according to a Poisson process with rate µ₀(u), u = 1,..., U.
- Customer type uniquely determines route through the network along the sequence of queues

 $r(u, 1), r(u, 2), \ldots, r(u, L(u)).$

- Customer may visit the same queue at multiple stages.
- Queue *j* operates according to the $(\kappa_j, \gamma_j, \delta_j)$ -protocol.
- Let c_j(ℓ) = (u_j(ℓ), s_j(ℓ)), with u_j(ℓ) the type and s_j(ℓ) the stage of the customer in position ℓ in queue j.
- State of queue *j* is $\mathbf{c}_j = (c_j(1), \dots, c_j(n_j))$.
- State of the network is $\mathbf{c} = (\mathbf{c}_1, \dots, \mathbf{c}_J)$.

Networks: customer types and fixed routes - 2

- ► Let {N(t)} record state of Markov chain at state space $S = {\mathbf{c} = (\mathbf{c}_1, ..., \mathbf{c}_J)}.$
- For $\mathbf{c} = (\mathbf{c}_1, \dots, \mathbf{c}_J)$, let
 - $C_{(\ell,j),(\ell',k)}^{(u,s)}$ **c** denote state **c**' obtained from state **c** by removing customer of type *u* in stage *s* in position ℓ from queue *j* and adding that customer in position ℓ' to queue *k*.
- Transition rates (more precise in reader)

$$\begin{aligned} q(\mathbf{c}, \mathbf{c}') &= \\ \begin{cases} \mu_0(u) \delta_k(\overline{\ell}', n_k + 1), & \text{if } \mathbf{c}' = C_{(0,0),(\ell',k)}^{(u,0)} \mathbf{c}, \\ \mu_j(u) \kappa_j(n_j) \gamma_j(\overline{\ell}, n_j) \delta_k(\overline{\ell}'_k, n_k + 1), & \text{if } \mathbf{c}' = C_{(\ell,j),(\ell',k)}^{(u,s)} \mathbf{c}, \\ \mu_j(u) \kappa_j(n_j) \gamma_j(\overline{\ell}, n_j), & \text{if } \mathbf{c}' = C_{(\ell,j),(0,0)}^{(u,0)} \mathbf{c}. \end{cases} \end{aligned}$$

• $\lambda_j(u, s)$: arrival rate of type *u* to queue j = r(u, s). Then

$$\lambda_j(u, s) = \begin{cases} \mu_0(u), & \text{if } j = r(u, s), \\ 0, & \text{otherwise.} \end{cases}$$

Mean amount of work arriving to queue j per unit time:

$$\rho_j = \sum_{u=1}^U \sum_{s=1}^{L(u)} \frac{\lambda_j(u,s)}{\mu_j(u)}, \quad j = 1, \dots, J.$$

• Let $\rho_j(c_j(\ell)) = \lambda_j(u_j(\ell), s_j(\ell))/\mu_j(u_j(\ell)).$

Theorem (4.3.1 Network with fixed routes)

Let queue *j* operate according to the $(\kappa_j, \gamma_j, \delta_j)$ -protocol. Negative-exponential(1) service requirements for all customers at all queues. Let

$$\pi_j(\mathbf{c}_j) = \mathbf{G}_j \prod_{\ell=1}^n \frac{\rho_j(\mathbf{c}_j(\ell))}{\kappa_j(\ell)}, \quad \mathbf{G}_j = \left[\sum_{n=0}^\infty \prod_{\ell=1}^n \frac{\rho_j}{\kappa_j(\ell)}\right]^{-1} < \infty,$$

Then

$$\pi(\mathbf{c}) = \prod_{j=1}^J \pi_j(\mathbf{c}_j), \quad \mathbf{c} \in S.$$

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Networks: customer types and fixed routes – 5

Proof. Natural guess for the reversed process:

- ► customers of type u arrive according to a Poisson process with rate µ₀(u) to queue L(u)
- ▶ and follow the reversed route $r(u, L(u)), \ldots, r(u, 1)$,
- and that the transition rates have the role of γ and δ reversed:

$$q^r(\mathbf{c}',\mathbf{c}) =$$

$$\begin{cases} \kappa_k(n_k+1)\delta_k(\overline{\ell}',n_k+1), & \text{if } \mathbf{c}' = C_{(0,0),(\ell',k)}^{(u,0)}\mathbf{c}, \\ \kappa_k(n_k+1)\delta_k(\overline{\ell}'_k,n_k+1)\gamma_j(\overline{\ell},n_j), & \text{if } \mathbf{c}' = C_{(\ell,j),(\ell',k)}^{(u,0)}\mathbf{c}, \\ \mu_0(u)\gamma_j(\overline{\ell},n_j), & \text{if } \mathbf{c}' = C_{(\ell,j),(0,0)}^{(u,0)}\mathbf{c}. \end{cases}$$

$$\sum_{\mathbf{c}'} q(\mathbf{c}, \mathbf{c}') = \sum_{u=1}^{U} \mu_0(u) + \sum_{j=1}^{J} \sum_{\ell_j=1}^{n_j} \kappa_j(n_j) \gamma_j(\ell_j, n_j),$$

$$\sum_{\mathbf{c}'} q^r(\mathbf{c}, \mathbf{c}') = \sum_{u=1}^{U} \mu_0(u) + \sum_{k=1}^{J} \sum_{\ell_k=1}^{n_k} \kappa_k(n_k) \delta_k(\ell_k, n_k),$$

For $\mathbf{c}' = C_{(\ell,j),(\ell',k)}^{(u,s)}\mathbf{c}$, with $j, k \neq 0$, we have $\pi(\mathbf{c})q(\mathbf{c}, \mathbf{c}') = \pi(\mathbf{c})\kappa_j(n_j)\gamma_j(\overline{\ell}, n_j)\delta_k(\overline{\ell}'_k, n_k + 1),$ $\pi(\mathbf{c}')q^r(\mathbf{c}', \mathbf{c})$ $= \pi(\mathbf{c})\frac{\rho_k(\mathbf{c}_k(\ell'_k))}{\rho_j(\mathbf{c}_j(\ell_j))}\frac{\kappa_j(n_j)}{\kappa_k(n_k + 1)}\kappa_k(n_k + 1)\delta_k(\overline{\ell}'_k, n_k + 1)\gamma_j(\overline{\ell}, n_j).$



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Quasi-reversibility - 1

- Burke's theorem: output process from a reversible queue before *t*, the input process after *t* and the state at *t* independent.
- Quasi-reversibility formalises this independence property.
- ► {N(t), $t \in \mathbb{R}$ } Markov process, state space *S*, states $\mathbf{n} \in S$, transition rates $q(\mathbf{n}, \mathbf{n}')$, equilibrium distribution $\pi(\mathbf{n})$.
- Let S(c, n) ⊂ S denote the set of states that may be obtained from state n when a customer of class c arrives to the queue.
- Let {A_c(t), t ∈ ℝ} and {D_c(t), t ∈ ℝ} record the arrival and departure processes of customers of class c.

Definition (4.4.1 Quasi-reversibility)

The stationary Markov chain $\{N(t)\}$ is quasi-reversible if for all $t \in \mathbb{R}$ the state at time t, N(t), is independent of $\{A_c(s), s > t\}$, the arrival process of class c customers after time t, and independent of $\{D_c(s), s < t\}$, the departure process of class c customers prior to time t, c = 1, ..., C.

Theorem (4.4.2)

If $\{N(t)\}$ is a quasi-reversible Markov chain, then

- (i) the arrival processes $\{A_c(t), t \in \mathbb{R}\}, c = 1, ..., C$, form independent Poisson processes;
- (ii) the departure processes $\{D_c(t), t \in \mathbb{R}\}, c = 1, ..., C$, form independent Poisson processes.

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Quasi-reversibility – 3

Algebraic characterisation of quasi-reversibility:

$$\lambda(c) = \sum_{\mathbf{n}' \in S(c,\mathbf{n})} q(\mathbf{n},\mathbf{n}'),$$

 $\lambda(c) = \sum_{\mathbf{n}' \in S(c,\mathbf{n})} q^r(\mathbf{n},\mathbf{n}'),$

so that

$$\sum_{\mathbf{n}'\in S(c,\mathbf{n})}\pi(\mathbf{n})q(\mathbf{n},\mathbf{n}')=\sum_{\mathbf{n}'\in S(c,\mathbf{n})}\pi(\mathbf{n}')q(\mathbf{n}',\mathbf{n}).$$

In equilibrium the flow out of state n due to a customer of type c arriving to the queue balances with the probability flow into state n due to a customer of type c departing from the queue.

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Symmetric queue - 1

Definition (4.2.6 Symmetric queue)

A queue that operates under the (κ, γ, δ) -protocol is called symmetric if

$$\gamma(\ell, n) = \delta(\ell, n), \quad \ell = 1, \dots, n, \ n \in \mathbb{N}.$$

Theorem (4.4.6)

Let {*N*(*t*)} record the state of a symmetric queue to which customers of class *c* arrive according to independent Poisson processes with rate $\lambda(c)$, c = 1, ..., C. Then {*N*(*t*)} is quasi-reversible.

Symmetric queue - 2

Proof.

► Transition rates, for $\mathbf{c} = (c(1), \dots, c(n)), \mathbf{c}' \neq \mathbf{c}$,

$$q(\mathbf{c},\mathbf{c}') = \begin{cases} \lambda(c)\gamma(\ell,n+1), & \text{if } \mathbf{c}' = (c(1),\ldots,c(\ell),c,c(\ell+1),\ldots,c(n)), \\ \mu_{c(\ell)}\kappa(n)\gamma(\ell,n), & \text{if } \mathbf{c}' = c(1),\ldots,c(\ell-1),c(\ell+1),\ldots,c(n)). \end{cases}$$

- ► Arrivals of class *c* customers independent Poisson processes ⇒ *N*(*t*) independent of {*A_c*(*s*), *s* > *t*}.
- Transition rates of time-reversed queue: $q^r = q$.
- Arrival process to the time-reversed queue is Poisson process.
- Arrivals in the time-reversed process coincide with departures of {*N*(*t*)} ⇒ *N*(*t*) is independent of {*D_c*(*s*), *s* < *t*}.



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Quasi-reversible queues and fixed routes - 1

- ► Network of *J* quasi-reversible queues.
- ► Customers of types u = 1,..., U, arrive to a according to a Poisson process with rate µ₀(u), u = 1,..., U.
- ► Customer type uniquely determines route along the sequence of queues r(u, 1), r(u, 2), ..., r(u, L(u)).
- ► State of queue *j*: {*N_j*(*t*)}, state space *S_j*, transition rates *q_j*(**c**_{*j*}, **c**'_{*j*}), customers of class (*u*, *s*) arrive according to Poisson process with rate

$$\lambda_j(u, s) \sum_{\mathbf{c}'_j \in S_j((u, s), \mathbf{c}_j)} q_j(\mathbf{c}_j, \mathbf{c}'_j),$$

► Equilibrium distribution $\pi_j = (\pi_j(\mathbf{c}_j), \mathbf{c}_j \in S_j)$ satisfies

$$\sum_{\mathbf{c}_j' \in S_j(c,\mathbf{c}_j)} \pi_j(\mathbf{c}_j) q_j(\mathbf{c}_j,\mathbf{c}_j') = \sum_{\mathbf{c}_j' \in S_j(c,\mathbf{c}_j)} \pi_j(\mathbf{c}_j') q_j(\mathbf{c}_j',\mathbf{c}_j).$$

- ▶ For $c = (c_1, ..., c_J)$, and j, k = 0, ..., J, let
 - $C_{j,k}^{(u,s)}$ **c** denote the set of states **c**' obtained from state **c** by removing the customer of type *u* in stage *s* from queue *j* and adding that customer in stage *s* + 1 to queue *k*:

$$(C_{j,k}^{(u,s)}\mathbf{c})_i = \begin{cases} \{\mathbf{c}_i\}, & \text{if } i \neq j, k, \\ S_k((u,s+1), \mathbf{c}_k), & \text{if } i = k, \\ \{\mathbf{c}'_j \text{ s.t. } \mathbf{c}_j \in S_j((u,s), \mathbf{c}'_j)\}, & \text{if } i = j, \end{cases}$$

Quasi-reversible queues and fixed routes - 3

► Transition rates, for u = 1, ..., U, $\mathbf{c} \neq \mathbf{c}', \mathbf{c}, \mathbf{c}' \in S$, $q(\mathbf{c}, \mathbf{c}') =$

$$\begin{cases} q_{k}(\mathbf{c}_{k},\mathbf{c}_{k}'), & \text{if } \mathbf{c}' \in C_{0,k}^{(u,1)}\mathbf{c}, & (\text{arrival}) \\ q_{j}(\mathbf{c}_{j},\mathbf{c}_{j}') \frac{q_{k}(\mathbf{c}_{k},\mathbf{c}_{k}')}{\sum_{\mathbf{c}_{k}' \in S_{k}((u,s+1),\mathbf{c}_{k})} q_{k}(\mathbf{c}_{k},\mathbf{c}_{k}')}, & \text{if } \mathbf{c}' \in C_{j,k}^{(u,s)}\mathbf{c}, & (\text{routing}) \\ q_{j}(\mathbf{c}_{j},\mathbf{c}_{j}'), & \text{if } \mathbf{c}' \in C_{j,0}^{(u,L(u))}\mathbf{c}, & (\text{departure}) \\ q_{j}(\mathbf{c}_{j},\mathbf{c}_{j}'), & \text{if } \mathbf{c}_{j},\mathbf{c}_{j}' \in S_{j}, \mathbf{c}_{i}' = \mathbf{c}_{i}, i \neq j, & (\text{internal}) \end{cases}$$

Quasi-reversibility implies that

$$\frac{q_k(\mathbf{c}_k,\mathbf{c}'_k)}{\sum_{\mathbf{c}'_k\in S_k((u,s+1),\mathbf{c}_k)}q_k(\mathbf{c}_k,\mathbf{c}'_k)} = \frac{q_k(\mathbf{c}_k,\mathbf{c}'_k)}{\lambda_k(u,s+1)}.$$

Theorem (4.5.1)

Let $\{N(t)\} = \{(N_1(t), ..., N_J(t))\}$ record the state of a network of J quasi-reversible queues to which customers of types u = 1, ..., U arrive according to independent Poisson processes with rates $\mu_0(u)$ to follow a fixed route r(u, 1), r(u, 2), ..., r(u, L(u)), u = 1, ..., U. Let S_j , q_j , and π_j denote the state space, transition rates and equilibrium distribution of queue j, j = 1, ..., J. Then $\{N(t)\}$ has equilibrium distribution

$$\pi(\mathbf{c}_1,\ldots,\mathbf{c}_J)=\prod_{j=1}^J\pi_j(\mathbf{c}_j),\quad (\mathbf{c}_1,\ldots,\mathbf{c}_J)\in S=S_1 imes\cdots imes S_J.$$

Quasi-reversible queues and fixed routes - 4

Proof. Natural guess for time-reversed process:

- ► customers of types u = 1,..., U arrive according to a Poisson process with rate µ₀(u),
- ► route through the network along the sequence of queues in reversed order r(u, L(u)), ..., r(u, 1)
- ► each queue operates according to its time-reversed transition rates: for u = 1,..., U, c ≠ c', c, c' ∈ S,

 $q^{\prime}(\mathbf{c}^{\prime},\mathbf{c}) =$

$$\begin{cases} q_k^r(\mathbf{c}'_k, \mathbf{c}_k), & \text{if } \mathbf{c}' \in C_{0,k}^{(u,1)}\mathbf{c}, & (\text{departure}) \\ q_k^r(\mathbf{c}'_k, \mathbf{c}_k) \frac{q_j^r(\mathbf{c}'_j, \mathbf{c}_j)}{\lambda_j(u, s)}, & \text{if } \mathbf{c}' \in C_{j,k}^{(u,s)}\mathbf{c}, & (\text{routing}) \\ q_j^r(\mathbf{c}'_j, \mathbf{c}_j), & \text{if } \mathbf{c}' \in C_{j,0}^{(u,L(u))}\mathbf{c}, & (\text{arrival}) \\ q_j^r(\mathbf{c}'_j, \mathbf{c}_j), & \text{if } \mathbf{c}_j, \mathbf{c}'_j \in S_j, \ \mathbf{c}'_i = \mathbf{c}_i, \ i \neq j. & (\text{internal}) \end{cases}$$

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For a routing transition from queue j = r(u, s) to queue k = r(u, s + 1) it must be that λ_j(u, s) = λ_k(u, s + 1), which implies that

$$\pi_j(\mathbf{c}_j)\pi_k(\mathbf{c}_k)q_j(\mathbf{c}_j,\mathbf{c}_j')\frac{q_k(\mathbf{c}_k,\mathbf{c}_k')}{\lambda_k(u,s+1)} = \pi^r(\mathbf{c}_j')\pi_k^r(\mathbf{c}_k')q_k^r(\mathbf{c}_k',\mathbf{c}_k)\frac{q_j^r(\mathbf{c}_j',\mathbf{c}_j)}{\lambda_j(u,s)}$$

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